

Introduction to Lower Previsions II

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http://gregorywheeler.org/courses/prolog/lp_2.pdf

Two ways of defining lower previsions w/out credal sets

Directly

- ▶ Define $\underline{P}(f)$ directly as Your supremum acceptable buying price for f
- ▶ $\underline{P}(f)$ is the highest price s s.t. for any $t < s$, You accept to pay t before observing X on the promise that You will receive $f(x)$ upon observing the event $X = x$.
- ▶ *Thus*, You are only required to consider whether you accept bounded gambles of type $f - \alpha$.
- ▶ *Conditioning is complicated*

Via Gambles

- ▶ Announce Your acceptable bounded gambles
- ▶ Enforce coherence conditions
- ▶ *Then* infer Your lower $\underline{P}(f)$ and upper $\overline{P}(f)$ previsions for any bounded gamble f .
- ▶ *Conditioning is clean(er)*

Three goals

1. *Get rid of probabilities, credal sets, and lower previsions*
 - ✓ Sets of **Acceptable Gambles**
Williams 1977; Walley 1991, 2000; De Cooman and Troffaes 2014
2. *Propose a benchmark for evaluating a model of belief*
 - ✓ **Operationalizable**: one-sided betting
 - ✓ **Inference**: closure, marginalization, and conditionalization
 - ✓ **Unification**: unify different uncertainty models
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3. *Address with some alleged problems*
 - **Dilation**
 - **Violations of Good's Principle**
 - No Strictly Proper IP Scoring Rules
 - ▶ Seidenfeld, Schervish, and Kadane (2012)
Drop quantifiability \Rightarrow *lexicographic decision theory*
 - ▶ Mayo-Wilson and Wheeler (forthcoming)
Drop strictness \Rightarrow *'normal' decision theory*

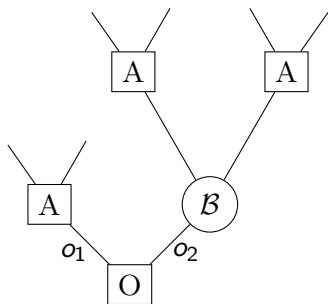
Outline

- Good's Principle
- Dilation Example
- Dilation Diagnoses
- Dilation Characterization Results (joint w/ Paul Pedersen)
- Plurality of Independence Concepts
- Rethinking Good's Principle

Good's Principle

Never turn down the offer of free information.

Good's Principle



You should delay making a terminal decision (A) between alternative courses of action if the opportunity arises to learn, at zero cost to you, the outcome of an experiment (B) relevant to the decision

Dilation – informal

Idea: A probability estimate in E becomes less precise upon learning the outcome of an experiment, \mathcal{B} , **no matter how the experiment turns out.**

Defining dilation

DEFN. Let $(\Omega, \mathcal{A}, \mathbb{P}, \underline{P})$ be a lower probability space, let \mathcal{B} be a positive measurable partition of Ω , and let E be an event.

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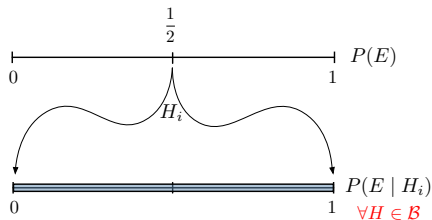
$$\underline{P}(E | H) < \underline{P}(E) \leq \bar{P}(E) < \bar{P}(E | H).$$

Defining dilation

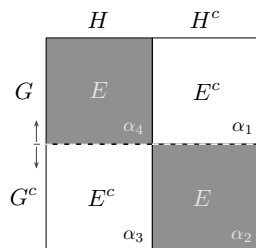
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EXAMPLE 1



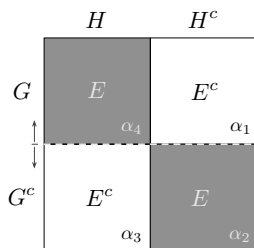
Dilation example



$\mathcal{B} = \{H, H^c\}$ outcomes of a fair coin toss,

$$\underline{P}(H) = \overline{P}(H) = \frac{1}{2} = \underline{P}(H^c) = \overline{P}(H^c) \quad (1)$$

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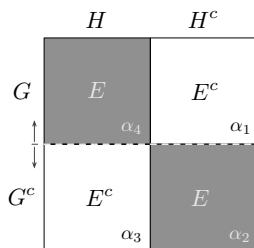
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$$\overline{P}(G) - \underline{P}(G) = .9 \quad (2)$$

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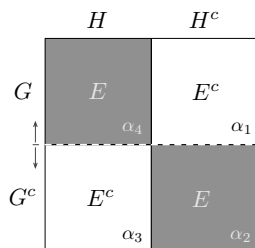
$$\overline{P}(G) - \underline{P}(G) = .9 \quad (2)$$

H and G are stochastically independent:

$$p(G \cap H) = p(G)p(H) = \frac{p(G)}{2}, \quad (3)$$

for each $p \in \mathbb{P}$.

Dilation example

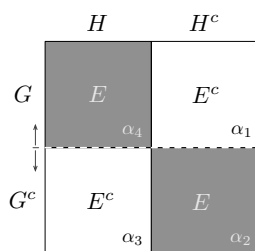


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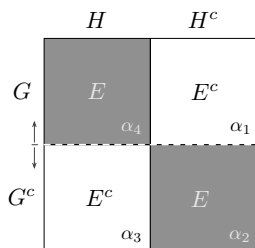
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DEFN. $E := (G \cap H) \cup (G^c \cap H^c)$

$$\text{So, } \forall p \in \mathbb{P}, p(E) = 1/2. \quad (4)$$

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Proof of (4):

$$\begin{aligned} \forall p \in \mathbb{P}: p(E) &= p(G \cap H) + p(G^c \cap H^c) \quad (\text{Defn. of } E) \\ &= \frac{p(G)}{2} + \frac{1 - p(G)}{2} \quad (\text{By Eq 3}) \\ &= \frac{p(G) + 1 - p(G)}{2} = \frac{1}{2}. \end{aligned}$$

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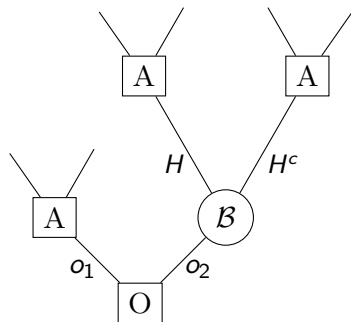
Observe: E is strictly dilated by the coin toss, $\mathcal{B} = \{H, H^c\}$.

Proof - We show that $0 \approx \underline{P}(E | H) < \underline{P}(E) = 1/2$.

$$\begin{aligned}\underline{P}(E | H) &= \inf \{ p(E | H) : p \in \mathbb{P} \} && \text{(Defn.)} \\ &= \inf \left\{ \frac{p([(G \cap H) \cup (G^c \cap H^c)] \cap H)}{p(H)} : p \in \mathbb{P} \right\} \\ &= \inf \left\{ \frac{p(G \cap H)}{p(H)} : p \in \mathbb{P} \right\} \\ &= \inf \left\{ \frac{p(G)p(H)}{p(H)} : p \in \mathbb{P} \right\} && \text{(By Eq 3)} \\ &= \inf \{ p(G) : p \in \mathbb{P} \} = .1.\end{aligned}$$

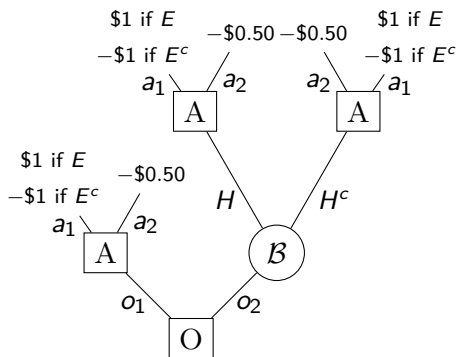
An analogous argument establishes $.9 = \overline{P}(E | H) > 1/2$; Finally, by the same reasoning $.1 = \underline{P}(E | H^c) < 1/2 < \overline{P}(E | H^c) = .9$. \diamond

Good's Principle

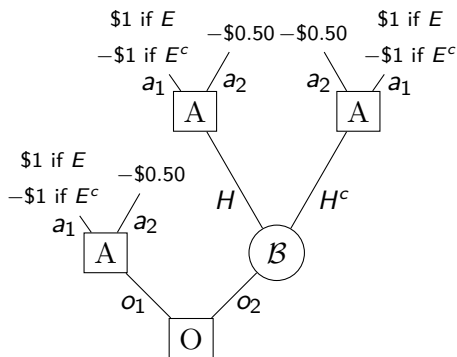


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Good's principle & minimum expected utility



Good's principle & minimum expected utility



$$o_1 : A = \begin{cases} 0 & \text{if } a_1 \\ -\$0.50 & \text{if } a_2 \end{cases}$$

$$o_2 : A | B = \begin{cases} -\$0.80 & \text{if } a_1^{**} \\ -\$0.50 & \text{if } a_2 \end{cases}$$

** for **some** $p \in \mathbb{P}$, i.e., $p(E | H^c) = .9$

Reactions to dilation

Dilation is pathological

- a) Reductio of imprecise probability theory
(Elga, White, J. Williamson)
- b) Spike Bayesian conditionalization
(Halpern & Grünwald, Kyburg)

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Dilation is benign

- c) Dilation is partially benign
(Joyce, Seidenfeld)
- d) Dilation is (completely?) benign
(Pedersen & Wheeler)

Two questions to address

What is dilation?

How does it figure in failures of Good's principle?

What accounts for dilation?

FACT 1 (Seidenfeld and Wassermann 1993)

Distance from independence is a necessary condition for dilation.

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DEFN. A simple measure of stochastic independence:

$$S_p(E, F) \quad =_{df} \quad \frac{p(E \cap F)}{p(E)p(F)}.$$

NOTE. $p(E | F) = p(E) \Rightarrow p(F | E) = p(F) \Rightarrow p(E \cap F) = p(E)p(F)$

What accounts for dilation?

FACTS (Seidenfeld and Wassermann 1993, 1994)

- If \mathcal{B} dilates E , any p realizing $\underline{P}(E | H)$ (any $p: \overline{P}(E | H)$) must make E and H negatively (positively) correlated (for every $H \in \mathcal{B}$).

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- If every H in \mathcal{B} is such that $\underline{P}(E)$ ($\overline{P}(E)$) is realized by some p for which E and H are negatively (pos.) correlated, then \mathcal{B} dilates E .
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Is there a general characterization of dilation?

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IDEA: A probability function is an element of a lower neighborhood of E conditional on H with radius ϵ if $p(E|H)$ is within ϵ of $\underline{P}(E|H)$

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DEFN. (Recall) Given a probability function p , a simple measure of stochastic independence:

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$p \in S_{\mathbb{P}}^+(E, F) \iff E$ and F are pos. assoc. with respect to p

$p \in S_{\mathbb{P}}^-(E, F) \iff E$ and F are neg. assoc. with respect to p

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That is:

$$S_{\mathbb{P}}^+(E, F) \quad =_{df} \quad \{p \in \mathbb{P} : S_p(E, F) > 1\}$$

$$S_{\mathbb{P}}^-(E, F) \quad =_{df} \quad \{p \in \mathbb{P} : S_p(E, F) < 1\}$$

$$I_{\mathbb{P}}(E, F) \quad =_{df} \quad \{p \in \mathbb{P} : S_p(E, F) = 1\}.$$

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THM. (Pedersen and Wheeler 2014)

Let $(\Omega, \mathcal{A}, \mathbb{P}, \underline{P})$ be a lower probability space such that \mathbb{P} is weak*-closed and convex,

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- (i) \mathcal{B} dilates E ;
- (ii) There is $(\epsilon_i)_{i \in I} \in \mathbb{R}_+^I$ such that for every $i \in I$:

$$\underline{\mathbb{P}}(E|H_i, \epsilon_i) \subseteq S^-(E, H_i) \quad \text{and} \quad \overline{\mathbb{P}}(E|H_i, \epsilon_i) \subseteq S^+(E, H_i);$$

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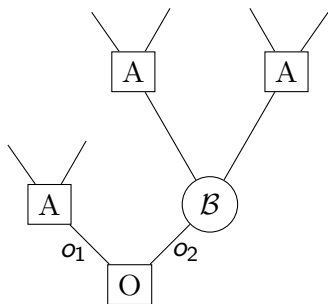
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Each radius $\underline{\epsilon}_i$ may be chosen to be the unique positive min of $|\rho(E | H_i) - \underline{P}(E | H_i)|$ on $C_i^+ := \{\rho \in \mathbb{P} : S_\rho(E, H_i) \geq 1\}$, unique $\overline{\epsilon}_i$ with respect to C_i^- .

Good's Principle



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Review of probabilistic independence

Two events E and F in \mathcal{A} are stochastically independent just in case:

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Two events E and F are epistemically independent just in case:

(EI) E is epistemically irrelevant to F and F is epistemically irrelevant to E .

IP independence

Our three familiar independence concepts:

(ER) Given $\underline{P}(F), \underline{P}(F^c) > 0$, event F is *epistemically irrelevant* to E if and only if:

1. $\underline{P}(E | F) = \underline{P}(E | F^c) = \underline{P}(E)$;
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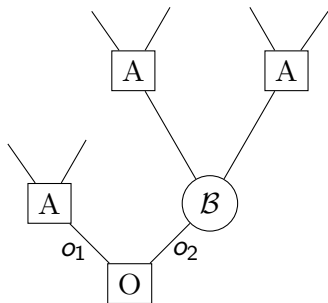
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(EI) E is *epistemically independent* of F just when both F is epistemically irrelevant to E and E is epistemically irrelevant to F .

(SI) E and F are *completely stochastically independent* if and only if for all $p \in \mathbb{P}$, $p(E \cap F) = p(E)p(F)$.

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Decision rules in IP

Assumptions in our delayed decision problem:

- ▶ You don't change your preferences
- ▶ You update by Generalized Bayesian conditionalization
- ▶ You *choose to maximize your minimum expected loss* (Γ -Maximin)

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- Reflection on strict Bayesian methods

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DEFN. A probability function is a real-valued function p defined on an algebra \mathcal{A} over a set of states Ω satisfying the following three conditions:

(P1) $p(E) \geq 0$ for every $E \in \mathcal{A}$;

(P2) $p(\Omega) = 1$;

(P3) $p(E \cup F) = p(E) + p(F)$ (\forall pairwise disjoint elements $E, F \in \mathcal{A}$).

(P4.1) $\underline{P}(E) = \inf \{p(E) : p \in \mathbb{P}\}$.

(P5.1) $\overline{P}(E) = \sup \{p(E) : p \in \mathbb{P}\}$.

(P4.2) $\underline{P}(E | F) = \inf \{p(E | F) : p \in \mathbb{P}\}$.

(P5.2) $\overline{P}(E | F) = \sup \{p(E | F) : p \in \mathbb{P}\}$.

(P6) If $\underline{P} = \overline{P}$, then $\{p\} = \mathbb{P}$ and $p = \underline{P} = \overline{P}$.

(P7) $\overline{P}(E) = 1 - \underline{P}(E^c)$.

(P8) $\underline{P}(E \cup F) \geq \underline{P}(E) + \underline{P}(F) - \underline{P}(E \cap F)$.

▶ return

Proof of PEDERSEN AND WHEELER, THM 5.1

We first show that (i) \iff (iii), and we then show that (i) \iff (ii).

(i) \implies (iii) Suppose that \mathcal{B} dilates E . Then for each $i \in I$,
 $\underline{P}(E|H_i) < \underline{P}(E) \leq \overline{P}(E) < \overline{P}(E|H_i)$. For each $i \in I$, consider
the real-valued function $\underline{\epsilon}_i(p) =_{df} |p(E|H_i) - \underline{P}(E|H_i)|$ and the
real-valued function $S_{p,i}(E, H_i)$. We recall that the weak*
topology on the dual space is a locally convex topological
vector space with respect to which every evaluation functional
 f^* is a real-valued continuous linear functional on the dual
space. It follows that $\underline{\epsilon}_i(p)$ and $S_{p,i}(E, H_i)$ are continuous
functions of p on \mathbb{P} for each $i \in I$. ▶ return (continued ...)

Now let $i \in I$. By hypothesis, there is $p_1 \in \mathbb{P}$ such that $S_{p_1, i}(E, H_i) > 1$, so importantly, C_i^+ is nonempty. Then since $C_i^+ =_{df} \{p \in \mathbb{P} : S_p(E, H_i) \geq 1\}$ is a weak*-closed and so weak*-compact set, it follows that $\underline{\epsilon}_i$ achieves a minimum value on C_i^+ (and the set of minimizers of $\underline{\epsilon}_i$ is also compact). Choosing a minimizer $p_i \in \mathbb{P}$ of $\underline{\epsilon}_i$, we see that for every $p \in \mathbb{P}$, if $|p(E|H_i) - \underline{P}(E|H_i)| < \underline{\epsilon}_i(p_i) = |p_i(E|H_i) - \underline{P}(E|H_i)|$, then $S_p(E, H_i) < 1$. We have accordingly shown that $\overline{\mathbb{P}}(E|H_i, \underline{\epsilon}_i(p_i)) \subseteq S_{\mathbb{P}}^-(E, H_i)$.

Of course, we may suppress reference to the minimizer p_i in $\underline{\epsilon}_i(p_i)$. The other inclusion $\overline{\mathbb{P}}(E|H_i, \bar{\epsilon}_i) \subseteq S_{\mathbb{P}}^+(E, H_i)$ is established by a similar argument. ▶ return (continued ...)

(iii) \Leftarrow (i) Suppose that \mathcal{B} does not dilate E . Then there is $i \in I$ such that $\underline{P}(E|H_i) \geq \underline{P}(E)$ or $\overline{P}(E) \geq \overline{P}(E|H_i)$. We may assume without loss of generality that $\underline{P}(E|H_i) \geq \underline{P}(E)$ for some $i \in I$. First, if $\underline{P}(E) \leq \overline{P}(E) \leq \underline{P}(E|H_i)$, then choosing a minimizer $p \in \mathbb{P}$ of $\underline{P}(E|H_i)$, we see that $S_p(E, H_i) \geq 1$. Second, if $\underline{P}(E) < \underline{P}(E|H_i) < \overline{P}(E)$, then for every $\epsilon > 0$ we can find a convex combination $p \in \mathbb{P}$ of $p_0, p_1 \in \mathbb{P}$ assigning a probability to E within ϵ -distance below $\underline{P}(E|H_i)$, where $\underline{P}(E) \leq p_0(E) < \underline{P}(E|H_i) < p_1(E) \leq \overline{P}(E)$, so $S_p(E, H_i) > 1$. Third, if $\underline{P}(E) = \underline{P}(E|H_i) < \overline{P}(E)$, then choosing a minimizer $p \in \mathbb{P}$ of $\underline{P}(E)$, we see that $S_p(E, H_i) \geq 1$. Evidently, the conditions of the main claim cannot be jointly satisfied.

[▶ return](#) (continued ...)

(i) \Leftrightarrow (ii) On the one hand, suppose that (i) obtains. Then since (iii) accordingly obtains, define $(\epsilon_i)_{i \in I}$ by setting $\epsilon_i =_{df} \min(\underline{\epsilon}_i, \bar{\epsilon}_i)$ for each $i \in I$. Clearly the inclusions still obtain for the ϵ_i . On the other hand, if (ii) obtains, obviously by setting $\underline{\epsilon}_i =_{df} \epsilon_i$ and $\bar{\epsilon}_i =_{df} \epsilon_i$ for each $i \in I$, condition (iii) obtains and so (i) obtains. \square

» return

2. Review of results

Here, we briefly review the results from Seidenfeld and Wasserman (1993). Given M , define $M_*(A) = \{P \in M; P(A) = \underline{P}(A)\}$, $M^*(A) = \{P \in M; P(A) = \bar{P}(A)\}$, $M_*(A|B) = \{P \in M; P(A|B) = \underline{P}(A|B)\}$ and $M^*(A|B) = \{P \in M; P(A|B) = \bar{P}(A|B)\}$. For $P \in \mathcal{P}$, the set of all probabilities, define $S_P(A, B) = P(A \cap B) / (P(A)P(B))$, $d_P(A, B) = P(A \cap B) - P(A)P(B)$, $\Sigma^+(A, B) = \{P \in \mathcal{P}; d_P(A, B) > 0\}$ and $\Sigma^-(A, B) = \{P \in \mathcal{P}; d_P(A, B) < 0\}$. The surface of independence for events A and B is defined by $\mathcal{I}(A, B) = \{P \in \mathcal{P}; d_P(A, B) = 0\}$. Finally, define

$$S_0 = \inf_{P \in M_*(A|B)} S_P(A, B), \quad S^0 = \sup_{P \in M^*(A|B)} S_P(A, B),$$

$$M_0 = \{P \in M_*(A|B); S_P = S_0\}, \quad M^0 = \{P \in M^*(A|B); S_P = S_0\},$$

and

$$P_0 = \inf_{P \in M_0} P(A), \quad P^0 = \sup_{P \in M^0} P(A).$$

Result 1. Suppose that $\mathcal{B} = \{B, B^c\}$. A necessary and sufficient condition for strict dilation is that

$$S_0 < \frac{P(A)}{P_0(A)} \leq 1 \leq \frac{\bar{P}(A)}{P^0(A)} < S^0.$$

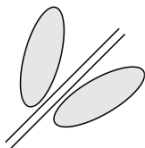
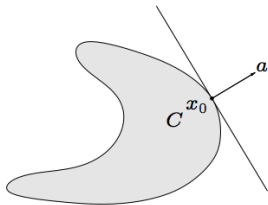


Figure 5.3. Strong separation.



Figure 5.4. These sets cannot be separated by a hyperplane.



▶ return